Economics Bulletin

Volume 32, Issue 3

Determinant Factors of Income Inequality – Evidence from a Portuguese Region

Nuno Crespo Instituto Universitário de Lisboa (ISCTE - IUL), UNIDE - IUL (Business Research Unit)

Nádia Simões Instituto Universitário de Lisboa (ISCTE - IUL), UNIDE - IUL (Business Research Unit) Ana Paula Diogo Banco de Portugal

Abstract

We propose a new methodology, based on bilateral comparisons between households, to analyze the determinants of income inequality, and confirm the influence of variables related to: (i) socio-economic characteristics of the household, (ii) dimension and composition of the household, (iii) the number of households in the residence, and (iv) socio-economic characteristics of the individual of reference.

The authors are grateful to an anonymous referee for helpful comments, to the Office of National Statistics (INE) for kindly providing us with the survey data, and to Fundação para a Ciência e para a Tecnologia for the financial support (PEst-OE/EGE/UI0315/2011). The usual disclaimer applies.

Citation: Nuno Crespo and Nádia Simões and Ana Paula Diogo, (2012) "Determinant Factors of Income Inequality – Evidence from a Portuguese Region", *Economics Bulletin*, Vol. 32 No. 3 pp. 2056-2064.

Contact: Nuno Crespo - nuno.crespo@iscte.pt, Nádia Simões - nadia.simoes@iscte.pt, Ana Paula Diogo - ana.paula.diogo@bportugal.pt. Submitted: May 14, 2012. Published: July 23, 2012.

1. Introduction

Income inequality is an important research area not only because of its impact on economic growth but also as support to economic policies that seek to reduce poverty and inequality (Fields, 2001). Research on this topic covers three main aspects: (i) methodological debate on how to measure inequality, (ii) quantification of the phenomenon in different countries/regions and periods, and (iii) identification of the determinants of inequality. In this study, we focus on this last feature, the least explored in economic literature.

The evaluation of the determinant factors of income inequality has been concretized through two main approaches: the traditional static inequality decomposition by income sources (Shorrocks, 1982) or by sub-groups (Shorrocks, 1984) and the regression-based inequality decomposition technique (Morduch and Sicular, 2002; Fields, 2003). The main purpose of this study is to propose a different methodology, based on a bilateral comparison of households, which allows a direct test through an econometric model of the determinant factors of income inequality.

The paper is organized as follows. Section 2 introduces the empirical methodology and presents the data. Section 3 summarizes the results. Section 4 concludes.

2. Empirical methodology

The traditional inequality decomposition approach proposed by Shorrocks (1982, 1984) has two important limitations. First, it has an eminently descriptive nature. Second, by construction the contribution(s) of the several characteristics of the household cannot be disentangled, and such knowledge may be important for economic policy (Naschold, 2009).

An important contribution to overcome these limitations is the regression-based inequality decomposition technique. This approach is a combination of income regression analysis and income source decomposition, allowing one to quantify how much inequality is explained by each income determinant.¹

In this study, we propose an alternative approach. An important distinctive characteristic of this method is the fact that, instead of explaining the overall level of income inequality, it uses micro-data to analyze, in bilateral terms, the determinants of inequality. This is done through an econometric model in which a measure of inequality between each pair of households is regressed as a function of the variables that explain the degree of similarity/dissimilarity between households. The number of observations in the model corresponds therefore to the total number of pairs that can be obtained from the data. Since the measure of income inequality is the dependent variable, it becomes possible to directly assess the impact of alterations in each explanatory variable on inequality, if, for a given pair of households, we have information about all the explanatory variables included in the model, we can predict the level of income inequality between them.

Let us now briefly discuss the methodological options that need to be assumed before empirical analyses on income inequality. Specifically, the following aspects must

¹ For examples of application of this methodology and discussion of their limitations, see, for instance, Wan and Zhou (2005) and Kimhi (2010).

be considered: (i) the indicator of resources, (ii) the demographic unit, and (iii) equivalence scales.

The indicator of resources most commonly used is the monetary disposable income, defined as the sum of work income, property income, pensions, other social transfers, and other private transfers after the deduction of the taxes on income and social contributions. However, this type of indicator excludes all forms of non-monetary income. In the current analysis, therefore, we use both monetary and total income. In this last case, the following items are also considered: the value of goods produced for own consumption, inputted rents, and remuneration in kind.

Regarding the demographic unit, following the widely accepted practice, we consider households.

A final issue that needs to be decided has to do with the comparison between unlike units. In fact, households with distinct compositions and dimensions require different incomes to achieve the same level of welfare. The use of equivalence scales allows calculating the household size in adult equivalents. In this study, we use the modified OECD scale, which gives a weight of 1 to the first adult, 0.5 to each of the remaining adults, and 0.3 for children under 14 years of age. When we divide the income (monetary and total) of each household by the number of equivalent adults, we obtain the adult equivalent income.

To measure income inequality between households *i* and *h*, we calculate:

$$INEQ_{i,h}^{\nu} = \begin{cases} \left| \frac{Y_i^{\nu}}{Y_i^{\nu} + Y_h^{\nu}} - \frac{Y_h^{\nu}}{Y_i^{\nu} + Y_h^{\nu}} \right| \text{ if } Y_i^{\nu} + Y_h^{\nu} > 0\\ 0 \text{ if } Y_i^{\nu} = Y_h^{\nu} = 0 \end{cases}$$
(1)

where $v = \{1, 2\}$, with 1 being monetary income, and 2 total income. Y expresses income per equivalent adult. *INEQ* is a relative measure of inequality and ranges between 0 (if the incomes of *i* and *h* are equal) and 1 (maximum inequality).

Regarding the way this variable is obtained, let us emphasize two aspects. First, since our focus is on the analysis of income inequality instead of poverty or richness, *INEQ* is built to attend to the magnitude of the gap between the two households in terms of income, ignoring therefore which of them is the richest. For this reason, we assume the absolute value of the difference between the relative weights of each household in the total income of the two households under consideration. Obviously, if we wish to extend the evaluation to the connected concepts of poverty and richness, we could consider the difference without taking the absolute value.

Second, through the consideration of the denominator $(Y_i^v + Y_h^v)$, we eliminate the sensibility of *INEQ* to the scale of measurement. A simple example illustrates this point. Let us consider four households (*A*, *B*, *C*, and *D*) with incomes $Y_a = 20$, $Y_b = 40$, $Y_c = 1000$, and $Y_d = 1020$. If we consider as dependent variable $|Y_i^v - Y_h^v|$, the level of inequality between the pairs *A-B* and *C-D* is the same. However, while *B* has 100% more income than *A*, the household *D* has only 2% more income than *C*, reflecting a much smaller degree of inequality.

Since *INEQ* measures the level of inequality between two households, our explanatory variables are constructed in order to capture the differences between these

households in the dimensions that are relevant to explain income inequality. We consider four groups of variables, related to: (i) socio-economic characteristics of the household (*ACT*, *UNEM*, *EDUC*, and *INC*), (ii) composition and dimension of the household (*INDV*, *ADU*, and *CHILD*), (iii) the number of households in the residence - e.g., single or multiple-family dwelling (*RES*), and (iv) socio-economic characteristics of the individual of reference (*LMS*, *GEND*, and *AGE*).² Table I presents the explanatory variables considered in the model.

Variable	Designation	Definition	
Active	ACT	Difference, in absolute value, between the ratios for each household, between the number of individuals in active age and the total number of individuals of the household	
Unemployment	UNEM	Difference, in absolute value, between the ratios for each household, between the number of unemployed and the total number of individuals in active age	
Education	EDU0	Dummy variable with value 1 if the average education level of the adults of the households is the same	
	EDU1	Dummy variable with value 1 if the difference, in absolute value, between the adults' average education level is equal to 1	
	EDU2	Dummy variable with value 1 if the difference, in absolute value, between the adults' average education level is equal to 2	
	EDU3	Dummy variable with value 1 if the difference, in absolute value, between the adults' average education level is equal to 3	
	EDU4	Dummy variable with value 1 if the difference, in absolute value, between the adults' average education level is equal to 4	
	EDU5	Dummy variable with value 1 if the difference, in absolute value, between the adults' average education level is equal to 5	
Income Source	INC	Dummy variable with value 1 if the main income sources of the households are different	
Individuals	INDV0	Dummy variable with value 1 if the number of individuals in each household is the same	
	INDV1	Dummy variable with value 1 if the difference, in absolute value, between the number of individuals in each household is equal to 1	
	INDV2	Dummy variable with value 1 if the difference, in absolute value, between the number of individuals in each household is equal to 2	
	INDV3	Dummy variable with value 1 if the difference, in absolute value, between the number of individuals in each household is equal to 3	
	INDV4	Dummy variable with value 1 if the difference, in absolute value, between the number of individuals in each household is equal to 4	
	INDV5	Dummy variable with value 1 if the difference, in absolute value, between the number of individuals in each household is equal to 5	
	INDV6	Dummy variable with value 1 if the difference, in absolute value, between the number of individuals in each household is equal to 6	
	INDV7	Dummy variable with value 1 if the difference, in absolute value, between the number of individuals in each household is equal to 7	
Adults	ADU	Dummy variable with value 1 if a household has only one adult and the other has more than one	
Children	CHILD	Dummy variable with value 1 if a household has children and the other does not	
Residence	RES	Dummy variable with value 1 if one of the households shares the	

Table I. Explanatory Variables

² The individual of reference of the household is the individual with the largest proportion of the annual net total income of the household.

		residence and the other does not
Labor Market State	LMS	Dummy variable with value 1 if the individuals of reference have different labor market states
Gender	GEND	Dummy variable with value 1 if the gender of the individual of reference is different
Age	AGE0	Dummy variable with value 1 if the individuals of reference belong to the same age group
	AGE1	Dummy variable with value 1 if the absolute value of the difference between the age groups of the individuals of reference is equal to 1
	AGE2	Dummy variable with value 1 if the absolute value of the difference between the age groups of the individuals of reference is equal to 2
	AGE3	Dummy variable with value 1 if the absolute value of the difference between the age groups of the individuals of reference is equal to 3
	AGE4	Dummy variable with value 1 if the absolute value of the difference between the age groups of the individuals of reference is equal to 4
	AGE5	Dummy variable with value 1 if the absolute value of the difference between the age groups of the individuals of reference is equal to 5

Notes: (i) the construction of the group of variables related to age (*AGE0* to *AGE5*) involves two steps. First, the age group of the individual of reference is identified: 1 (age between 18 and 24), 2 (25-29), 3 (30-44), 4 (45-64), 5 (65-74), and 6 (more than 74). Second, for each pair of households, we calculate the absolute value of the difference of the age groups of the individuals of reference; (ii) the construction of the dummy variables related to the comparison between the number of individuals (*INDV0* to *INDV7*) is similar to the previous one. We start by identifying the number of individuals in each household and then calculate the difference in absolute terms between these numbers; (iii) the dummy variables related to households' average education are obtained in three steps. First, the education level of each adult is identified: 1 (no education), 2 (primary education – 1^{st} cycle), 3 (primary education – 2^{nd} cycle), 4 (primary education – 3^{rd} cycle), 5 (secondary education), and 6 (tertiary education). Second, for each households, the difference, in absolute value, between these average education levels. This difference is then rounded to the nearest unit.

When the dependent variable is bounded, using OLS regression may produce biased and inconsistent parameter estimates. One of the models that can be used instead, overcoming these problems, is the Tobit regression model (Tobin, 1958). Since *INEQ* varies between 0 and 1, meaning that the dependent variable of our model is bounded below and above, the two-limit Tobit regression is adequate (Rosett and Nelson, 1975).

Based on earlier evidence reported in the literature on poverty and inequality, the expected sign of the variables presented in Table I is positive, except for the case of: (i) the variables concerning the dimension and composition of the household, for which there is no obvious expected sign, and (ii) the variables related with the age of the individual of reference. In the latter case, the difficulty stems from the fact that higher incomes are usually concentrated in households with middle age individuals of reference while households with younger and older individuals of reference have, on average, lower incomes.

We use micro-data from the Office of National Statistics (INE) Household Budget Survey (IDEF) carried out in 2005/2006. Our sample includes the households residing in the region of *Lisboa* (1317 households).

3. Results

Table II presents the results from the estimation of the model discussed above.

	Depender	nt variables
Explanatory variables	INEQ ¹ _{i,h}	$\mathbf{INEQ}_{_{i,h}}^{^{2}}$
ACT	0.019***	0.021***
	(13.35)	(14.90)
UNEM	0.028***	0.034***
EDUI	(24.32)	(30.22)
EDU1	0.012***	0.010***
EDU2	(12.73) 0.045***	(11.36)
EDU2		0.041***
EDU3	(48.00) 0.103***	(45.10) 0.097***
EDUS	(105.66)	(102.49)
EDU4	0.191***	0.180***
LDU4	(177.23)	(173.05)
EDU5	0.306***	0.304***
ED05	(189.46)	(194.74)
INC	0.015***	-0.009***
live	(24.95)	(-15.88)
INDV1	-0.002**	-0.003***
, .	(-2.86)	(-4.70)
INDV2	-0.001	-0.001**
11072	(-0.91)	(-2.02)
INDV3	0.0167***	0.0190***
11010	(16.57)	(19.58)
INDV4	0.0432***	0.0566***
	(27.02)	(36.59)
INDV5	0.039***	0.064***
	(14.08)	(23.77)
INDV6	0.058***	0.076***
	(12.63)	(17.19)
INDV7	0.049***	0.058***
	(5.18)	(6.34)
ADU	-0.003***	-0.003***
	(-5.05)	(-5.63)
CHILD	-0.018***	-0.019***
	(-30.07)	(-31.64)
RES	0.013***	0.014***
	(4.22)	(4.78)
LMS	0.002**	0.004***
	(2.50)	(6.48)
GEND	0.007***	0.006***
	(13.50)	(13.36)
AGE1	-0.005***	-0.002***
	(-8.00)	(-3.34)
AGE2	-0.011***	-0.005***
	(-15.73)	(-6.87)
AGE3	-0.033***	-0.023***
	(-33.95)	(-25.15)
AGE4	-0.048***	-0.035***
1055	(-25.72)	(-19.36)
AGE5	-0.072***	-0.055***
CONSTANT	(-20.56)	(-16.11)
CONSTANT	0.279***	0.277***
	(281.86)	(289.24)
Sigmo	0 214***	0.200***
Sigma	0.216***	0.209***
	(1316.25)	(1316.46)
Log-likelihood	96656.87	126071.28
Number of observations	866586	866586
	00000	00000

Table II. Determinant Factors of Income Inequality

Note: *, **, *** significant at 10%, 5% and 1%, respectively.

Regarding monetary income inequality, the results presented in Table II confirm the expected influence of variables related to the socio-economic characteristics of the household, i.e., the inequality increases: (i) with the difference between the number of individuals in active age in the total number of individuals of the households, (ii) with the difference between the number of unemployed individuals in the total of active age individuals, (iii) with the difference between the average educational level of the adults of the households, and (iv) if the main sources of income are different. Among these effects, differences in terms of education are especially important to explain inequality. The evidence obtained is in accordance with earlier studies which have shown the importance of education levels as an important determinant of the labor market situation (Card, 1999). Individuals with higher levels of education are more likely to hold jobs that involve performing more complex tasks and that have better career perspectives and retirement plans. Additionally, they are more efficient in job searching (Arrow, 1997).

The impact of the variables related with the dimension and composition of the households is the most difficult to identify *a priori*. The evidence reveals that significant differences between households concerning the number of individuals that compose them, particularly when the difference is greater than three, tend to increase inequality. On the other hand, however, the variables *ADU* and *CHILD* have a negative impact on the level of income inequality. We should note that this result is not inconsistent with the evidence obtained from *INDV* because, while *INDV* is capturing quantitative differences in the dimension of the households, *ADU* and *CHILD* aim to provide some insights regarding their composition.

The variable related to the number of households in the residence confirms the expected sign. Thus, inequality is higher when we compare households in which one lives alone and the other shares the residence than when the number of households in the residence is the same.

When we consider the variables related to the individual of reference, the evidence suggests that inequality is higher when the gender of the individual of reference is different. This conclusion is clearly in line with theoretical and empirical research regarding the existence of important differences between genders in the labor market concerning, for instance, pay levels, promotions, and distribution across occupations (Altonji and Blank, 1999; Blau and Kahn, 2006). Discrimination behaviors from employers and co-workers, productivity differences, and differences in preferences, namely concerning competitive environments, are possible explanations for the less favorable situation of women. All these aspects have potential implications in terms of income, as shown, for instance, by De Silva (2008) for developing countries.

Additionally, confirming the expected impact, income inequality increases when the labor market states of the individuals of reference are different, as also suggested by the conclusions obtained by Moller et al. (2003).

A greater difference between the age of the individuals of reference is associated with lower inequality. Two reasons may explain this result. First, considering the distribution of the average income by the age of the individuals of reference, there is an inverted U shaped curve, in which households with a younger or older individual of reference have very similar incomes. Second, the middle-age group is, by far, the one where income dispersion is the greatest.

The evidence based on total income confirms the results obtained with monetary income, the only exception being *INC*, which now has a negative coefficient, indicating that households with different main sources of income have a lower degree of income inequality.

4. Conclusion

In this study we proposed a new approach for assessing the determinants of inequality in income distribution. This methodology, which is based on a bilateral comparison of income, was used to analyze inequality in the Portuguese region of *Lisboa*. The results show that inequality increases as a function of differences in: (i) the socio-economic characteristics of households (active age individuals, number of unemployed individuals, the main source of income, and average educational levels), (ii) aspects related to the dimension and composition of the household (number of individuals, adults, and children), (iii) the number of households in the residence, and (iv) the socioeconomic characteristics of the individual of reference (gender, labor market state, and age). Considering all the variables, we also conclude that the largest impact on income inequality is caused by differences in terms of average education levels.

References

Altonji, J. and R. Blank (1999) "Race and gender in the labor market" in *Handbook of Labor Economics* (vol. 3C) by O. Ashenfelter and D. Card, Eds., Amsterdam: North-Holland, 3143-3259.

Arrow, K. (1997) "The benefits of education and the formation of preferences" in *The social benefits of education* by J. Behrman and N. Stacy, Eds., University of Michigan Press: Ann Arbor, 11-16.

Blau, F. and L. Kahn (2006) "The U.S. gender pay gap in the 1990s: slowing convergence" *Industrial & Labor Relations Review* **60**, 45-66.

Card, D. (1999) "The causal effect of education on earnings" in *Handbook of Labor Economics* (vol. 3A) by O. Ashenfelter and D. Card, Eds., Amsterdam: North-Holland, 1801-1863.

De Silva, I. (2008) "Micro-level determinants of poverty reduction in Sri Lanka: a multivariate approach" *International Journal of Social Economics* **35**, 140-158.

Fields, G. (2001) Distribution and development: a new look at the developing world, Russell Sage Foundation: New York.

Fields, G. (2003) "Accounting for income inequality and its change: a new method, with application to the distribution of earnings in the United States" *Research in Labour Economics* **22**, 1-38.

Kimhi, A. (2010) "Jewish households, Arab households, and income inequality in rural Israel: ramifications for the Israeli-Arab conflict" *Defence and Peace Economics* **21**, 381-394.

Moller, S., D. Bradley, E. Huber, F. Nielsen and J. Stephens (2003) "Determinants of relative poverty in advanced capitalist democracies" *American Sociological Review* **68**, 22-51.

Morduch, J. and T. Sicular (2002) "Rethinking inequality decomposition with evidence from rural China" *Economic Journal* **112**, 93-106.

Naschold, F. (2009) "Microeconomic determinants of income inequality in rural Pakistan" *Journal of Development Studies* **45**, 746-768.

Rossett, R. and F. Nelson (1975) "Estimation of the two-limit Tobit regression model" *Econometrica* **43**, 141-146.

Shorrocks, A. (1982) "Inequality decomposition by factor component" *Econometrica* **50**, 193-211.

Shorrocks, A. (1984) "Inequality decomposition by population subgroup" *Econometrica* **52**, 1369-1385.

Tobin, J. (1958) "Estimation of relationships for limited depended variables" *Econometrica* **26**, 24-36.

Wan, G. and Z. Zhou (2005) "Income inequality in rural China: regression-based decomposition using household data" *Review of Development Economics* **9**, 107-120.