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### Exchange Rate Dynamics and Fundamental Equilibrium Exchange Rates

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#### Abstract

The paper investigates if the most popular alternative to the purchasing parity power approach (PPP) to estimate equilibrium exchange rates, the fundamental equilibrium exchange rate (FEER) influences exchange rate dynamics in the long run. For a large panel of industrialized and emerging countries and on the period 1982-2007, we detect the presence of unit roots in the series of real effective exchange rates and in the series of FEERs. We find and estimate a cointegration relationship between real effective exchange rates and FEERs. The results show that the FEER has a positive and significant influence on exchange rate dynamics in the long run.

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## 1. Introduction

The evolution of current account balances, which have considerably increased since the last fifteen years, has raised many questions. Among them, many observers (see e.g. Cline and Williamson, 2010) think that some countries (mainly South East Asian countries) have pursued aggressive exchange rate policy to maintain their currency at artificially low level in order to preserve their competitiveness on foreign markets. These politics of exchange rate undervaluation have allowed these countries to accumulate huge current account surplus. However as the misalignments (deviation from equilibrium exchange rate) offset each other at the world scale, if some countries are undervalued and run large current account surplus then some other countries are overvalued and run large current account deficits.

The paper investigates if the most popular alternative to the PPP approach to estimate equilibrium exchange rates, the FEER (Williamson, 1983, 1994) influences exchange rate dynamics in the long run. If the FEER affects the long run exchange rate dynamics, it can be considered as the 'true equilibrium exchange rate' and policy makers should pursue this rate to ensure global macroeconomic stability in a context of large imbalances.

As pointed out by Clark and MacDonald (1998), the notion of equilibrium in the FEER, which is a method of calculation of the equilibrium exchange rate consistent with external equilibrium (sustainable capital flows) and internal equilibrium (full utilization of productive potential), can be questioned if we consider that some variables which influence directly exchange rate behavior are often omitted in the calculations.

Zhou (1993) tests if the FEER affects exchange dynamics thanks to unit root tests and bivariate cointegration tests. She found empirical evidence that real effective exchange rates (REERs) and FEERs are unit roots process however she does not found that REERs and FEERs are cointegrated and, then, she concluded that exchange rate dynamics is not affected by the FEER.

These results can be misleading because they concerned only two countries (Japan and Germany) and a relatively short span of time (1974-1988) whereas the FEER approach is essentially a multilateral approach in which the global consistency, mentioned above, is crucial.

Barisone et alii (2006) study this question by using data for the G7 countries on the period 1973-1997. They implemented unit root tests, panel unit root tests and panel cointegration tests (Pedroni, 1999). They found empirical evidence that REERs and FEERs are unit roots process and, contrary to Zhou (1993), they found that REERs and FEERs are cointegrated.

For a large panel of industrialized and emerging countries and on the period 1982-2007, we detect the presence of unit roots in the series of REERs and in the series of FEERs. We find and estimate a cointegration relationship between REERs and FEERs. The results show that the FEER has a positive and significant influence on exchange rate dynamics in the long run.

The paper is organized as follow. Section 2 presents panel unit root tests. Section 3 tests a long run relationship between REERs and FEERs by using Pedroni's panel cointegration tests (1999). Section 4 estimates the long run relationship between REERs and FEERs by using the fully modified ordinary least square estimator (FMOLS) and the dynamic ordinary least square estimator (DOLS) introduced by Pedroni (2001). We also use a Pooled Mean Group (PMG) estimator introduced by Pesaran et alii (1999). Section 5 deals with the implications of the results in terms of international monetary cooperation. Section 6 concludes.

## 2. Panel unit root tests for REERs and FEERs

In this section, we implement a number of panel unit root tests for the series of REERs<sup>1</sup> and for the series of FEERs of seventeen industrialized and emerging countries (the U.S., the U.K., the Euro area, Japan, Korea, China, Brazil, India, Mexico, Argentina, Chile, Colombia, Indonesia, Malaysia, Philippines, Thailand and Uruguay) and on the period 1982-2007.

The FEER is defined as the exchange rate prevailing when the economy simultaneously reaches the external equilibrium and the internal equilibrium for all the trading partners. This measure was derived from a standard world trade model in which all the variables are endogenous except the external equilibrium (sustainable current account determined by structural variables) and the internal equilibrium (full utilization of the productive potential). The external equilibrium is estimated with panel regression techniques. The internal equilibrium is reached when the output gap is closed<sup>2</sup>.

**Table 1: Panel unit root tests**

<i>Test:</i>	<i>LLC</i>	<i>Breit.</i>	<i>F_ADF</i>	<i>F_PP</i>	<i>LLC</i>	<i>Breit.</i>	<i>F_ADF</i>	<i>F_PP</i>
<i>Difference:</i>	No	No	No	No	Yes	Yes	Yes	Yes
<i>Exogenous variable:</i>	Constant, Trend	Constant, Trend	Constant, Trend	Constant, Trend	None	None	None	None
<i>Null Hypothesis:</i>	UR	UR	UR	UR	UR	UR	UR	UR
<i>Common UR:</i>	Yes	Yes	No	No	Yes	Yes	No	No
<i>reer<sub>i,t</sub></i>	0.5	2.7	27.3	27.9	-15.1***	-11.9***	267.4***	286.0***
<i>feer<sub>i,t</sub></i>	-1.1	1.4	38.4	62.2***	-17.4***	-13.7***	312.8***	363.9***

Notes: "UR" indicates the null hypothesis of the presence of unit root. The symbol \*\*\* indicates statistical stationarity at the 1 percent level. The table shows different panel unit root tests: Levin, Lin, and Chu (2002) (LLC); Breitung (2000); Maddala and Wu (1999) and Choi (2001) Fischer-type panel unit root tests (F\_ADF and F\_PP).

Source: authors' calculations.

As we can see in table 1, we detect the presence of unit root in the series of REERs and in the series of FEERs. The series are nonstationary in level and stationary in first difference. The REERs and FEERs are nonstationary I(1) series. As a series is I(1) if it achieves stationarity after first differencing. This result is confirmed by other empirical studies (Zhou, 1993, Barisone et alii, 2006).

<sup>1</sup> Source: Bank for International Settlements for the real effective exchange rate basis 100 in 2000 (annual average of monthly data).

<sup>2</sup> See Jeong et al. (2010), Afrouk et al. (2010). The methodology used is a synthesis of previous works on the FEER (Borowski and Couharde, (2003), Jeong and Mazier, (2003)) and of the Symmetric Matrix Inversion Method (SMIM) recently proposed by Cline (2008).

### 3. Testing a long run relationship between REERs and FEERs

In order to test a long run relationship between REERs and FEERs, we can conduct either panel unit root tests on the difference between REERs and FEERs or panel cointegration tests (Pedroni, 1999).

As we have found that the series of REERs and the series of FEERs are nonstationary then if we found that the difference between these two series is stationary, it would mean that these series are cointegrated with a (1; -1) coefficient. In this paper, we estimate the long run relationship between REERs and FEERs thanks to the FMOLS estimator and the DOLS estimator introduced by Pedroni (2001) so it is unnecessary to impose any value to the cointegration vector before the empirical estimation (Barisone et alii, 2006). As a robustness check, we use the Pooled Mean Group estimator PMG (Pesaran et alii, 1999). We test the following long run equation<sup>3</sup>:

$$reer_{i,t} = \alpha_i + \beta feer_{i,t} + \mu_{i,t} \quad (1)$$

where variables in minuscule represents natural logarithms. As we can see in table 2, the null hypothesis of no cointegration is strongly rejected in all of the tests. These results show that there is a long run relationship between REERs and FEERs.

**Table 2: Panel cointegration tests<sup>4</sup>**

<i>Pedroni residual cointegration test (1999)</i>	
<i>Null Hypothesis: No cointegration</i>	
<i>Included observations</i>	442
<i>Cross-sections included</i>	17
<i>Alternative hypothesis: common AR coefficients (within-dimension)</i>	
<i>Panel-v</i>	2.51*
<i>Panel-rho</i>	-3.01***
<i>Panel-PP</i>	-3.65***
<i>Panel-ADF</i>	-3.98***
<i>Alternative hypothesis: individual AR coefficients (between-dimension)</i>	
<i>Group rho-Statistic</i>	-0.99
<i>Group PP-Statistic</i>	-3.22***
<i>Group ADF-Statistic</i>	-5.82***

Notes: The symbols \*, \*\*\* indicate statistical significance at the 10 percent and 1 percent levels, respectively. \*\*\*Reject null of nonstationarity even at the 1% level; \*Reject null of nonstationarity at the 10% level.

Source: authors' calculations.

<sup>3</sup> We do not include temporal trend as we want to test if REERs and FEERs are cointegrated without any divergence (Barisone et alii, 2006). We use the natural logarithms of these variables in order to obtain directly the elasticities.

<sup>4</sup> In appendix A, we provide panel unit root tests (Pesaran, 2007) and panel cointegration tests (Westerlund, 2007) which allows for cross section dependencies (i.e. existence of common shocks).

Before to proceeding to the next step which will consist to estimate explicitly this long run relationship. We have to conduct panel cointegration test (Pedroni, 1999) between the FEER and the REER since, as it is described in appendix B, the causal relationship is bi-directional i.e. the FEER causes the REER and the REER causes the FEER. We test the following long run equation:

$$feer_{i,t} = \delta_i + \theta reer_{i,t} + \varepsilon_{i,t} \quad (2)$$

As we can see in table 3, the null hypothesis of no cointegration is strongly rejected in all of the tests. These results show that there is a long run relationship between FEERs and REERs.

**Table 3: Panel cointegration tests**

<i>Pedroni residual cointegration test (1999)</i>	
<i>Null Hypothesis: No cointegration</i>	
<i>Included observations</i>	442
<i>Cross-sections included</i>	17
<i>Alternative hypothesis: common AR coefficients (within-dimension)</i>	
<i>Panel-v</i>	0.85
<i>Panel-rho</i>	-2.50***
<i>Panel-PP</i>	-3.38***
<i>Panel-ADF</i>	-3.39***
<i>Alternative hypothesis: individual AR coefficients (between-dimension)</i>	
<i>Group rho-Statistic</i>	-1.59*
<i>Group PP-Statistic</i>	-4.06***
<i>Group ADF-Statistic</i>	-5.42***

Notes: The symbols \*, \*\*\* indicate statistical significance at the 10 percent and 1 percent levels, respectively. \*\*\*Reject null of nonstationarity even at the 1% level; \*Reject null of nonstationarity at the 10% level.

Source: authors' calculations.

#### 4. Estimation of the long run relationship between REERs and FEERs

In this section, we implement the FMOLS estimator, the DOLS estimator and the PMG estimator to estimate the long run relationship between REERs and FEERs, as it is specified in equation (1) and equation (2). We use these econometric tools, introduced by Pedroni (2001) and Pesaran et alii (1999), in a context of nonstationarity of the series to avoid problems of fallacious regressions.

The results, in table 4 to 7, indicate that the FEER has a positive and significant influence on exchange rate dynamics in the long run<sup>5</sup>. In order to check the sensitivity to the results at the type of countries included in the panel, we re-estimate the equation (1) and equation (2) for emerging countries only, the results are largely similar.

<sup>5</sup> The FMOLS, DOLS and PMG estimations give similar results with a long run coefficient close to 0.6 for equation (1) and (2).

**Table 4: Long run relationship between REERs and FEERs**

	<i>Long Run Coefficient (<math>\beta</math>)</i>	<i>T-Statistic</i>
<i>FMOLS</i> <sup>1</sup>	0.66***	24.06
<i>DOLS</i> <sup>2</sup>	0.65***	29.37
<i>PMG</i> <sup>3</sup>	0.68***	11.05
<i>Hausman test</i>	0.15	0.70
<i>Cross-section included</i>	17	
<i>Number of Observations</i>	442	

Notes: (1) FMOLS is the Fully Modified OLS estimation; (2) DOLS is the Dynamic OLS estimation; (3) PMG is the Pooled Mean Group estimation. The symbol \*\*\* indicates statistical significance at the 1 percent level. The null hypothesis in the Hausman test is homogeneity of the long run coefficient in the PMG estimation.

Source: authors' calculations

**Table 5: Long run relationship between REERs and FEERs (emerging countries only)**

	<i>Long Run Coefficient (<math>\beta</math>)</i>	<i>T-Statistic</i>
<i>FMOLS</i> <sup>1</sup>	0.68***	22.28
<i>DOLS</i> <sup>2</sup>	0.66***	26.72
<i>PMG</i> <sup>3</sup>	0.66***	10.04
<i>Hausman test</i>	1.73	0.18
<i>Cross-section included</i>	13	
<i>Number of Observations</i>	338	

Notes: (1) FMOLS is the Fully Modified OLS estimation; (2) DOLS is the Dynamic OLS estimation; (3) PMG is the Pooled Mean Group estimation. The symbol \*\*\* indicates statistical significance at the 1 percent level. The null hypothesis in the Hausman test is homogeneity of the long run coefficient in the PMG estimation.

Source: authors' calculations

**Table 6: Long run relationship between FEERs and REERs**

	<i>Long Run Coefficient (<math>\theta</math>)</i>	<i>T-Statistic</i>
<i>FMOLS</i> <sup>1</sup>	0.62***	23.26
<i>DOLS</i> <sup>2</sup>	0.62***	27.64
<i>PMG</i> <sup>3</sup>	0.65***	15.36
<i>Hausman test</i>	1.43	0.23
<i>Cross-section included</i>	17	
<i>Number of Observations</i>	442	

Notes: (1) FMOLS is the Fully Modified OLS estimation; (2) DOLS is the Dynamic OLS estimation; (3) PMG is the Pooled Mean Group estimation. The symbol \*\*\* indicates statistical significance at the 1 percent level. The null hypothesis in the Hausman test is homogeneity of the long run coefficient in the PMG estimation.

Source: authors' calculations

**Table 7: Long run relationship between FEERs and REERs (emerging countries only)**

	<i>Long Run Coefficient (<math>\theta</math>)</i>	<i>T-Statistic</i>
<i>FMOLS</i> <sup>1</sup>	0.66***	22.09
<i>DOLS</i> <sup>2</sup>	0.66***	25.69
<i>PMG</i> <sup>3</sup>	0.71***	16.34
<i>Hausman test</i>	0.85	0.35
<i>Cross-section included</i>	13	
<i>Number of Observations</i>	338	

Notes: (1) FMOLS is the Fully Modified OLS estimation; (2) DOLS is the Dynamic OLS estimation; (3) PMG is the Pooled Mean Group estimation. The symbol \*\*\* indicates statistical significance at the 1 percent level. The null hypothesis in the Hausman test is homogeneity of the long run coefficient in the PMG estimation.

Source: authors' calculations

## 5. Implication of the results in terms of international monetary cooperation

After the crisis, we have observed a reduction of global imbalances mainly due to production decrease without currencies realignments. In 2011, the exchange rate misalignments remain important in the main world economies as noted by Cline and Williamson (2011). The United-States remains overvalued by around 7 % in real effective terms. China is undervalued by 17.6% in real effective terms and 28.5% in bilateral terms against the dollar. The Euro area is close to equilibrium and Japan is slightly overvalued in real effective terms and undervalued by around 10% in bilateral nominal terms<sup>6</sup>. For different reasons, the Euro Area (Euro crisis) and Japan (economic consequences of the Earthquake and Tsunami) are more constrained by their exchange rate policy than China.

Gagnon (2011) argues that the global imbalances should increase substantially (in larger proportions than the IMF's predictions) in the next years if we do not undertake the appropriate measures to prevent this widening. Feldstein (2011) discuss on the role of the currencies realignments to eliminate global imbalances. He insists on the complementary of domestic policies (private consumption or government spending) and external policies (currencies realignments) to cope with the global imbalances and ensure global macroeconomic stability.

Zhou (1993) concludes that if FEERs are not related to spot rates, either in the short run or in the long run, why they should viewed as "equilibrium exchange rate" and why government should intervene on foreign exchange rate markets or establish "target zones" for exchange rates based on them? Her empirical results can be misleading because they concern only two countries and relatively short span of time whereas the FEER approach is essentially a multilateral approach in which the global consistency, mentioned above, is crucial.

Our results showed that the FEER affects exchange rate dynamics for a large panel of countries and on the period 1982-2007. These results ensure that the FEER approach can be a useful tool for policy makers to compute currencies realignments in an international monetary cooperation which aims to eliminate global imbalances.

## 6. Conclusion

We provide empirical evidence that the most popular alternative to the purchasing parity power approach (PPP) to estimate equilibrium exchange rates, the fundamental equilibrium exchange rate (FEER) influences exchange rate dynamics in the long run. For a large panel of industrialized and emerging countries and on the period 1982-2007, we detect the presence of unit roots in the series of real effective exchange rates and in the series of FEERs. We find and estimate a cointegration relationship between real effective exchange rates and FEERs.

The results show that the FEER has a positive and significant influence on exchange rate dynamics in the long run. Therefore it can be considered as the 'true equilibrium exchange rate' and policy makers should pursue this rate to ensure global macroeconomic stability in a context of large imbalances.

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<sup>6</sup> Japan is undervalued in bilateral nominal terms against the dollar but as its East Asian competitors are relatively more undervalued in bilateral nominal terms against the dollar, Japan is slightly overvalued in real effective terms.

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### Appendix A: Panel unit root tests and panel cointegration tests

We use the CADF test introduced by Pesaran (2007) to test the unit root properties of the variables in presence of cross section dependencies. By subtracting cross section averages of lagged levels in addition to the standard ADF equation, this test is robust to cross section dependencies. The series are nonstationary I(1) series.

**Table A1: Integration of the variables involved**

	Level	First Difference
<i>feer</i>	0.223 (0.806)	-4.271*** (0.000)
<i>reer</i>	0.375 (0.646)	-3.083*** (0.001)

Notes: The p-values are in parentheses. The symbol \*\*\* indicates statistical stationarity at the 1 percent level.  
Source: authors' calculations.

To test cointegration, the panel and group mean statistic suggested by Westerlund (2007) are applied. The existence of a negative and significant error correction term is taken as proof for cointegration. In case of cross section dependencies between members of the panel, critical values need to be obtained through bootstrapping. Results indicate that these variables are cointegrated.

**Table A2: Cointegration of the variables involved**

	$G\tau$	$G\alpha$	$P\tau$	$P\alpha$
<i>reer, feer</i>	-2.162*** (0.001)	-6.414** (0.013)	-6.661* (0.070)	-4.005** (0.031)
<i>feer, reer</i>	-2.481*** (0.000)	-7.460*** (0.000)	-9.441*** (0.001)	-6.548*** (0.001)

Notes: The symbols \*, \*\*, \*\*\* indicate statistical significance at the 10 percent, the 5 percent and 1 percent levels, respectively. The p-value for cointegration tests are based on bootstrap methods, where 800 replications are used. See Persyn and Westerlund (2008) for the details.  
Source: authors' calculations.

### Appendix B: Panel Causality Tests

We test the following relationship:

$$reer_{i,t} = \alpha_{0i} + \alpha_{1i} feer_{i,t} + u_{i,t} \quad (3)$$

where variables in minuscule represents natural logarithms. The ARDL<sup>7</sup> equation (1, 1) associated with the above equation can be written:

$$reer_{i,t} = \delta_{0i} + \delta_{1i} feer_{i,t} + \delta_{2i} feer_{i,t-1} + \lambda_i reer_{i,t-1} \quad (4)$$

We can rewrite equation (4) as follows:

$$\Delta reer_{i,t} = \lambda_i reer_{i,t-1} - reer_{i,t-1} + \delta_{0i} + (\delta_{1i} + \delta_{2i}) feer_{i,t} - \delta_{2i} \Delta feer_{i,t} + \varepsilon_{i,t} \quad (5)$$

The error correction equation yield:

$$\Delta reer_{i,t} = \phi_i (reer_{i,t-1} - \theta_{0i} - \theta_{1i} feer_{i,t}) - \delta_{2i} \Delta feer_{i,t} + \varepsilon_{i,t} \quad (6)$$

with  $\phi_i = -(1 - \lambda_i)$ ;  $\theta_{0i} = \delta_{0i} / (1 - \lambda_i)$ ;  $\theta_{1i} = (\delta_{1i} + \delta_{2i}) / (1 - \lambda_i)$

Since the PMG estimator imposes long-run coefficients to be constant for all individuals, while it allows short run heterogeneity, the error correction model is written:

$$\Delta reer_{i,t} = \phi (reer_{i,t-1} - \theta_0 - \theta_1 feer_{i,t}) - \delta_{2i} \Delta feer_{i,t} + \varepsilon_{i,t} \quad (7)$$

$$\Delta feer_{i,t} = \phi (feer_{i,t-1} - \theta_0 - \theta_1 reer_{i,t}) - \delta_{2i} \Delta reer_{i,t} + \varepsilon_{i,t} \quad (8)$$

We estimate two reciprocal equations for causality test with the PMG estimator (i.e. we test if the REER causes the FEER and if the FEER causes the REER). The decision rule is: when the error correction term is negative and significant then the null hypothesis of no causality is rejected.

**Table B1: Panel Causality Tests**

	<i>Error-corrections coefficient (<math>\phi</math>)</i>
<i>FEER</i> → <i>REER</i>	-0.24*** (-7.30)
<i>REER</i> → <i>FEER</i>	-0.38*** (-6.37)
<i>Number of Observations</i>	442

Notes: t-statistics in parentheses: \*\*\* significant at 1 percent.  
Source: authors' calculations.

<sup>7</sup> Autoregressive Distributed Lags.

**Table B2: Panel Causality Tests (emerging countries only)**

	<i>Error-corrections coefficient (<math>\phi</math>)</i>
<i>FEER</i> → <i>REER</i>	-0.27*** (-7.74)
<i>REER</i> → <i>FEER</i>	-0.42*** (-6.05)
<i>Number of Observations</i>	338

Notes: t-statistics in parentheses: \*\*\* significant at 1 percent.  
Source: authors' calculations.

We can conclude that:

- The no causality hypothesis is strongly rejected in all cases
- The error-correction coefficients ( $\phi$ 's) are negative and statistically significant, indicating the causal relationship is bi-directional.
- These relationships are robust to the types of countries.

On the one side, the first kind of causality (equation (7)) may reveal cyclical evolution of competitiveness since the REER returns to its equilibrium (i.e. the FEER) in the long run. For example, if a country depreciates its currency in order to cope with competitiveness problems after some periods the REER return to the FEER if the country does not undertake specific measures to improve its competitiveness in the long run. On the other side, the second kind of causality (equation (8)) may express structural evolution of competitiveness since the equilibrium (i.e. the FEER) moves in order to conserve a constant, stable relationship with the REER (i.e. they are cointegrated). If a country experienced an appreciation of its currency reflecting a structural improvement of its competitiveness, the FEER will also appreciate. In other words, the country can compete in foreign markets with higher price thanks to the structural improvement of its competitiveness.

In case of cyclical evolution of competitiveness, the half-life<sup>8</sup> is equal to 3.2 years (2.9 years for emerging countries only). For structural evolution of competitiveness, the half-life is equal to 2.1 years (1.9 years for emerging countries only). When a country experienced a cyclical evolution of its competitiveness, it can slow the return to equilibrium in case of unfavorable evolutions hence a longer half-life.

<sup>8</sup> The half-lives are computed by using the following formula:  $h = -\ln(0.5)/\ln(1+|\phi|)$ . They correspond to the number of periods for a deviation (from the long run equilibrium) to decay by 50%.